| Introduction | Health shock process | lifecycle model | calibration | results |
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The Lifetime Costs of Bad Health

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Workshop on the Economics of Ageing

De Nardi, Pashchenko, and Porapakkarm

Lifetime Cost of Bad Health

#A. Large difference in economic outcomes by health

Among men with high-school degree, on average ...

- i. The healthy earn 37% more (conditional on working)...
- ii. ...and have 65 % more wealth at the time of retirement

▶ Wealth gradient (HRS)

#B. Two important questions

- What generates this large difference?
- How costly it is to be unhealthy from the entire life-cycle perspective?

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| Linking hea | Ith and economic ou | itcomes | | |

- Ch.1: Health affects economic outcomes
- *Ch.2:* Economic outcomes affect health
- Ch.3: Healthy and unhealthy people are innately different
- \Rightarrow *Ch.3* is well-recognized but overlooked (or too simplified) in existing structural studies
- \Rightarrow Our paper combines *Ch.1* with detailed investigation of *Ch.3*

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| Innate diffe | rences between the h | nealthy and unhea | althy | |

What is *Ch.3*?

- People differ in genetics, personality traits, early life experiences, etc.
- Growing empirical literature emphasizes the importance of these factors for outcomes later in life. (Anda et al., 2006; Barth et al., 2020; Case et al., 2005; Conti et al.,2005; among many others)
- We introduce these complex unobserved heterogeneity into a structural life-cycle model
- People differ in fixed characteristics that are multi-dimensional and possibly correlated among each other.

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| What we do? | The broad pictu | re | | |

1st **Part** : Estimate health shock process

- New data facts related to duration dependence of health status
- Formulate and estimate heath shock process that is consistent with these facts
- ► Key Finding :
 - *Health types* are an important driver of health dynamics even controlling for long history-dependence

| Introduction | Health shock process | lifecycle model | calibration | results |
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| What we do? | The broad pictur | e (cont.) | | |

2nd **Part**: Study interaction of health and economic outcomes in a structural model

Estimate a life cycle model augmented with the health shock and correlated ex-ante heterogeneity:

{ health type, fixed labor productivity, patience }

- Show that the correlated heterogeneity is important in explaining disparity in economic outcomes by health
- Quantify how costly it is to be unhealthy

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| Data | | | | |

- 1. Health and Retirement Study (HRS: 1994-2016)
- 2. Panel Study of Income Dynamics (PSID)
 - Annual data (1984-1997); bi-annual (1997-2017)
- 3. Medical Expenditure Panel Survel (MEPS: 1999-2017)

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| Outline of | the presentation | | | |



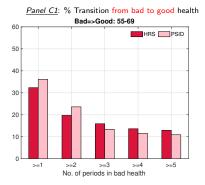
► Life-cycle model

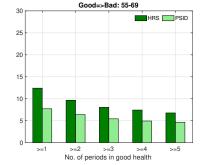
► Model estimation (MSM)



Introduction Health shock process of occurrence of the shock process of

Duration-dependent profile by health status (55-69 years old)





Panel C2: % Transition from good to bad health

- The difference between waves is 2 years



| Introduction | Health shock process | lifecycle model | calibration | results |
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| Health show | ck process | | | |

How we can account for these facts?

- Duration dependence
- Fixed health type
- Heterogeneity within bad health state

Formulate ordered logit model of health shock that allows for

- History-dependence (τ_B, τ_G) and discrete health type (η)
- Different transitions probabilities for two subcategories of bad health (B): fair (F) and poor (P)

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| Health sho | ock process | | | |

If $h_t \in \{P, F\}$ and duration of bad health (*P* or *F*) is τ_B :

$$logit\left[Pr\left(P_{t+1} \mid h_{t}, \tau_{B}, \eta\right)\right] = \underbrace{f_{age}^{h}\left(t\right)}_{age \text{ polynomial}} + \underbrace{\sum_{\tau=1}^{T-1} a_{\tau}^{B} \mathbf{1}_{(\tau_{B}=\tau)} + a_{T}^{B} \mathbf{1}_{(\tau_{B}\geq T)}}_{duration \text{ dependence}} + \underbrace{a_{\eta}^{B} \mathbf{D}_{\eta}}_{health \text{ type}}$$
$$logit\left[Pr\left(F_{t+1} \cup P_{t+1} \mid h_{t}, \tau_{B}, \eta\right)\right] = f_{age}^{h}\left(t\right) + \sum_{\tau=1}^{T-1} a_{\tau}^{B} \mathbf{1}_{(\tau_{B}=\tau)} + a_{T}^{B} \mathbf{1}_{(\tau_{B}\geq T)} + b_{1} + a_{\eta}^{B} \mathbf{D}_{\eta}$$

• If $h_t = G$ and duration of good health is τ_G :

$$logit\left[Pr\left(P_{t+1} \mid G_{t}, \tau_{G}, \eta\right)\right] = f_{age}^{G}\left(t\right) + \sum_{\tau=1}^{T-1} a_{\tau}^{G} \mathbf{1}_{(\tau_{G}=\tau)} + a_{T}^{G} \mathbf{1}_{(\tau_{G}\geq T)} + a_{\eta}^{G} \mathbf{D}_{\eta}$$
$$logit\left[Pr\left(F_{t+1} \cup P_{t+1} \mid G_{t}, \tau_{G}, \eta\right)\right] = f_{age}^{G}\left(t\right) + \sum_{\tau=1}^{T-1} a_{\tau}^{G} \mathbf{1}_{(\tau_{G}=\tau)} + a_{T}^{G} \mathbf{1}_{(\tau_{G}\geq T)} + b_{2} + a_{\eta}^{G} \mathbf{D}_{\eta}$$

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Lifetime Cost of Bad Health

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| Health sho | ock process (cont.) | | | |

Health type prediction

- η is distributed over 3 discrete points
- Ordered logit model of health type prediction:

$$logit \left[Pr(\eta_1 \mid \mathbf{X}_{t_0}) \right] = \mathbf{B}_{\eta} \mathbf{X}_{t_0}$$

$$logit \Big[Pr(\eta_1 \cup \eta_2 \mid \mathbf{X}_{t_0}) \Big] = \mathbf{B}_{\eta} \mathbf{X}_{t_0} + b_{\eta_2}$$

- $Pr(\eta_1 \cup \eta_2 \cup \eta_3 \mid \mathbf{X}_{t_0}) = 1$
- t_0 is the first age an individual was observed in the data.
- \mathbf{X}_{t_0} : initial health, initial wealth, fixed labor productivity (γ), age t_0 , birth cohort (10-year bracket)

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| Results : | Key findings | | | |

 Health type is always significant even when controlling for long lagged health history (up to 8 years)

• Health type (η) is correlated with fixed labor productivity (γ)

| | η_1 | η_2 | η_3 |
|------------------------------------|----------|----------|----------|
| $Pr(\eta)$ | 0.08 | 0.35 | 0.57 |
| $Pr(\eta \mid \gamma_L)$ | 0.13 | 0.44 | 0.43 |
| $Pr(\eta \mid \gamma_M)$ | 0.08 | 0.36 | 0.56 |
| $Pr\left(\eta\mid\gamma_{H} ight)$ | 0.04 | 0.24 | 0.72 |

Measure of η at age 21 (T=3)

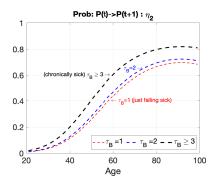
(* Use initial health, fixed labor productivity, wealth among people (21-24) in PSID)

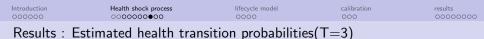
Health type prediction

model vs data

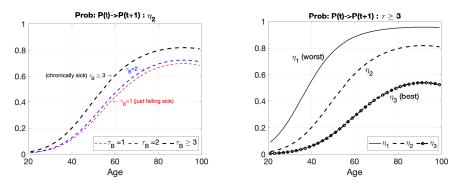
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| Results · | Estimated health trans | ition probabilities | (T=3) | |

History dependence : fix health type to η_2



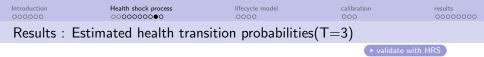


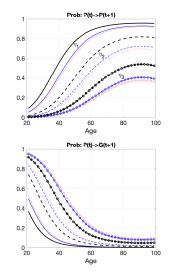
History dependence vs. Fixed health type

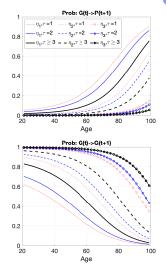


Key findings

Variation in health transition probabilities across health types is much larger than that across health histories







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How should we think about health type?

Model: People with bad health type experience multiple periods being unhealthy

► *HRS:* Characteristics of people by #periods being unhealthy

| # unhealthy | nealthy Individuals' characteristics ^a (HRS) | | | | $\% \eta_1$ | |
|-------------|---|------------------|----------------|----------------|----------------------------------|----------|
| periods | % smoking | BMI ^b | % father alive | % mother alive | parents' educ (yrs) ^c | in model |
| 0-1 | 22.6 | 27.9 | 21.6 | 48.4 | 10.1 / 10.5 | 0.1 |
| 2-3 | 27.1 | 29.5 | 21.5 | 50.4 | 9.2 / 9.9 | 3.1 |
| 4-5 | 44.4 | 29.8 | 16.1 | 36.5 | 8.4 / 9.2 | 26.0 |

^a All variables are reported at age 55-56.

^b BMI is the average Body Mass Index.

^c The first and second numbers are the average educational years of father and mother, respectively.

| # unhealthy | Polygenic scores (HRS) | | | | |
|-------------|------------------------|---------|--------|-----------|--|
| periods | educational attainment | smoking | BMI | longevity | |
| 0-1 | -0.120 | 0.003 | -0.006 | -0.06 | |
| 2-3 | -0.216 | 0.023 | 0.127 | -0.065 | |
| 4-5 | -0.708 | 0.092 | 0.140 | -0.250 | |

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Life-cycle model

► Model estimation (MSM)



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| Key mechar | nisms | | | |

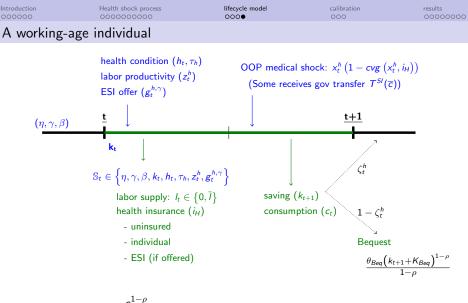
- The observed correlation between health and life-cycle outcomes is generated by two mechanisms
- 1 Causal effects of bad health:
 - a. Decreases productivity and increases disutility from work
 - b. Increases OOP medical spending
 - c. Lowers life expectancy
- 2 Composition effect:
 - Heterogeneity in health types (η), fixed productivity (γ), and patience (β)
 - $\{\eta, \gamma, \beta\}$ are correlated

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| Life-cycle r | nodel | | | |

- ▶ 21-64 \rightarrow work, 65-99 \rightarrow retired ...(model period = 2 yrs)
- Health type: $\eta \in \{\eta_1, \eta_2, \eta_3\}$ and discount factor: $\beta \in \{\beta_{low}, \beta_{high}\}$

 $0 \leq Pr(\beta_j | \eta_m) \leq 1; j \in \{low, high\}, m \in \{1, 2, 3\}$

- People face productivity, health, medical expenses, and survival uncertainty
- Retired people receive Social Security benefits and are covered by Medicare



$$u(c_t, l_t, h_t) = \frac{c_t}{1 - \rho} - \phi_W \mathbf{1}_{\{l_t > 0\}} - \phi_F \mathbf{1}_{\{h_t = F, l_t > 0\}} - \phi_P \mathbf{1}_{\{h_t = P, l_t > 0\}} + \overline{\mathbf{b}}$$

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► Life-cycle model

Model estimation (MSM)

Results

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Model parameters taken/estimated outside model

| parameters | | sources |
|------------------------------------|-------------------|-------------------------------|
| Survival probability by health: | ζ_t^h | HRS |
| | | (extrapolation from 21 to 50) |
| Health transition probability: | | PSID |
| Labor productivity shock: | z_t^h | PSID |
| Health-dependent medical expenses: | x_t^h | MEPS |
| Insurance coverage: | $cvg(x_t^h, i_H)$ | MEPS |
| ESI offer probability (logit) : | $g_t^{h,\gamma}$ | MEPS |
| Risk aversion: | $\rho = 3.0$ | common values $\in [1, 5]$ |

Parameters taken/estimated outside model

Iabor productivity shock

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Parameters estimated inside model

| parameters | value | | | targets |
|--|-----------------------|------------------------|------------------------|---|
| $\{\beta_{low}, \beta_{high}\}$ | $\{0.877, 0.992\}$ | | | " |
| $Pr(eta_{low} \eta_i)$ | $\frac{\eta_1}{0.78}$ | η ₂ 0.79 | η ₃ 0.38 | net wealth profiles by health (PSID) |
| consumption floor (per year): \overline{c} | | \$3,505 | | 11 |

* η_1 is the worst health type

- $\overline{\mathbf{b}} \Rightarrow$ Statistical Value of Life (SVL)
 - Compensation for adding 1 death among 10,000 adults:
 - Empirical SVL = 1-16M USD
 - Model: average SVL among working-age individuals = 2M USD

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| Results | | | | |

- R1. The importance of compositional difference
- R2. Lifetime monetary losses due to bad health
- R3. Lifetime welfare losses due to bad health

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R1 : The importance of compositional difference

Re-estimate the model but restricting $Pr(\beta_{low}|\eta_i) = 0.50$.

Wealth difference between healthy and unhealthy people at ages 60-64.

| Wealth difference by health | PSID (HRS) | Baseline | No (β,η) correlation |
|-----------------------------|---------------|----------|-------------------------------|
| 25 th pct | \$56 (\$47) | \$67 | \$38 |
| 50 th pct | \$142 (\$98) | \$146 | \$38 |
| 75 th pct | \$210 (\$222) | \$260 | \$91 |
| in 1000USD | | | |

IN 100005D

- No correlation between types and patience misses health-wealth gradient

- Income-health gradient does not imply wealth-health gradient

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| R1. The m | onetary cost of bad | health | | |

- Construct "always healthy" counterfactual
- Individuals always draw good health (unexpectedly)
- Let y_t^{BS} and y_t^H are income net of total medical expenses in baseline and "always healthy" cases.
- Measure of lifetime monetary losses :

$$\frac{1}{T} \sum_{t=1}^{T} \frac{y_t^H - y_t^{BS}}{(1+r)^t}$$

T is age at death

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R2. Lifetime monetary losses due to bad health

| | Over entire life-cycle (21-death) | | | |
|--|-----------------------------------|-------------------------|------------------------|-----------------|
| | All | η_1 | η_2 | η_3 |
| % of time in bad health | 15% | 58% | 23% | 4% |
| Annual monetary losses (% of avg earning) | \$1,511 <i>(3.9%)</i> | \$8,896 <i>(23%)</i> | \$1,935 <i>(5%)</i> | \$225 (0.6%) |
| Composition (%) | | | | |
| - Medical losses paid by insurance | 36% | 33% | 39% | 39% |
| - Out-of-pocket medical losses | 27% | 22% | 30% | 36% |
| - Income losses | 37% | 45% | 31% | 24% |

- Using 2% interest rate

- Average earning (2013) is \$38,648

- \blacktriangleright Losses vary a lot across η
- Income losses account for almost 40%



R3. Lifetime welfare losses due to bad health

Again, construct "always healthy" counterfactual

Measure of lifetime welfare losses due to bad health

Individual's life time utility in the baseline and "always heathy" cases:

$$U^{BS} = \sum_{t=1}^{T_d+1} \beta^t \Big(u(\boldsymbol{c}_t^*, \boldsymbol{l}_t^*, \boldsymbol{h}_t) \times \boldsymbol{1}_{alive_t} + (1 - \boldsymbol{1}_{alive_t}) \theta_{Beq} \frac{(\boldsymbol{k}_t^* + \boldsymbol{k}_{Beq})^{1-\rho}}{1-\rho} \Big),$$

$$U^{G}\left(\lambda_{c}\right) = \sum_{t=1}^{T_{d}^{C}+1} \beta^{t} \left(u\left((1-\lambda_{c})\boldsymbol{c}_{t}^{**}, \boldsymbol{l}_{t}^{**}, \boldsymbol{h}_{t} = good\right) \times \boldsymbol{1}_{\textit{alive}_{t}} + (1-\boldsymbol{1}_{\textit{alive}_{t}}) \theta_{\textit{Beq}} \frac{\left(\boldsymbol{k}_{t}^{**} + \boldsymbol{k}_{\textit{Beq}}\right)^{1-\rho}}{1-\rho}\right)$$

• Lifetime welfare losses $= \lambda_c \overline{c}^{**}$ where $\rightarrow U^{BS} = U^G (\lambda_c)$ $\rightarrow \overline{c}^{**}$ is the average consumption in "always healthy" case

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| R3. Lifetim | e welfare losses | | | |

| | all | η_1 | η_2 | η_3 |
|---|---------|----------|----------|----------|
| Compensated consumption equivalence | \$1,933 | \$6,380 | \$2,690 | \$854 |
| (% consumption equivalence, λ_c) | (10.6%) | (36.8%) | (14.8%) | (4.4%) |
| Contribution (%) | 050/ | 220/ | 221/ | |
| Only medical expenses channel | 25% | 39% | 22% | 17% |
| Only income channel | 38% | 57% | 42% | 9% |
| - Only survival channel | 44% | 32% | 33% | 77% |

Using SVL=\$2M.

- \blacktriangleright Welfare losses vary a lot across η
- On average, survival channel is the most important channel for welfare loss
- Income channel is the most important for {η₁, η₂} while the survival channel is the most important for η₃.

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R3. Lifetime losses due to bad health: concentration and contribution of η

| | Concentration | | | variation | |
|--|---------------|---------|---------|---------------|--|
| - | top 5% | top 10% | top 20% | due to η | |
| Monetary losses (21-death) | | | | | |
| - Income losses + medical losses (Ins+OOP) | 38% | 56 % | 75% | 69% | |
| Welfare losses | | | | | |
| - Compensated consumption equivalence | 24% | 42% | 71% | 30% | |

Use 2% interest rate for monetary loss.

Highly concentrated

 $\blacktriangleright\,$ A large variation in both monetary and welfare loss is due to $\eta\,$

\blacktriangleright But the variation due to η is lower for welfare losses

- η directly affects the number of periods being unhealthy
- But η only indirectly affects life expectancy.

variation in T^d

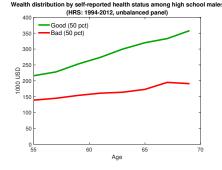
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Conclusions

- We quantify the effects of health in a life-cycle model of high school males that matches
 - (1) Long-run health dynamics
 - (2) Income-health gradient
 - (3) Wealth-health gradient
- Health type: important to capture (1)
- Composition difference btw. the healthy and unhealthy: important for (3)
- We measure lifetime loss due to bad health
 - i. Lifetime costs of bad health are highly concentrated
 - ii. Survival channel attributes a lot to welfare loss
 - iii A large variation in lifetime losses are pre-determined in early stage of life (69% for monetary loss, 30% for welfare loss)

Is the accumulated loss due to bad health large?

Wealth-health gradient among high school men (HRS: 1994-2016)



- good health \in {excellent, very good, good}; bad health \in {fair, poor}

- net worth: controlled for year effects and family sizes

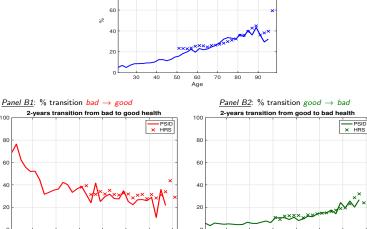
The wealth gap is large even among a relatively homogeneous group

Bias measure of losses if there is a composition difference

▶ back

Health status data (PSID, HRS)





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Aae

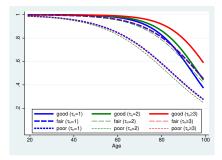


Aae

Estimated health-dependent survival probability (HRS: 1994-2017)

Logit regression of survival probability

$$logit\left(\zeta_{t}|h_{t},\tau_{h}\right) = \begin{cases} \underbrace{f_{age}^{\zeta^{h}}}_{age \text{ polynomial}} + \sum_{\tau=1}^{2} a_{\tau}^{\zeta B} \mathbf{1}_{(\tau_{h}=\tau)} + a_{3}^{\zeta B} \mathbf{1}_{(\tau_{h}=3)} & \text{if } h_{t} \in \{P,F\} \\ \\ f_{age}^{\zeta^{G}} + \sum_{\tau=1}^{2} a_{\tau}^{\zeta G} \mathbf{1}_{(\tau_{h}=\tau)} + a_{3}^{\zeta G} \mathbf{1}_{(\tau_{h}=3)} & \text{if } h_{t} = G. \end{cases}$$



Predicting health type

What observables (\mathbf{X}_{t_0}) are informative about health type (η) ?

▶ Initial health (h_{t_0})

Fixed labor productivity (γ)

Fixed effect regression of log labor income $log(inc_{it}) = \sum_{age=20}^{65} \sum_{j=\{G,B\}} d_t^j \cdot D_{it}^{age} \cdot \mathbf{D}_{h_{it}=j} + \gamma_i + u_{it},$ • FE estimation

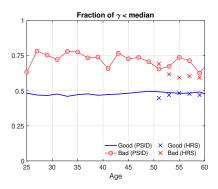
lntitial net worth (k_{t_0})

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Health status and fixed productivity γ

Fixed effect regression of log labor income :

$$log(inc_{it}) = \sum_{age=20}^{65} \sum_{j=\{G,B\}} d_t^j \cdot D_{it}^{age} \cdot \mathbf{D}_{h_{it}=j} + \gamma_i + u_{it},$$

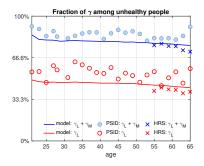


more low fixed productivity among unhealthy

Health status and fixed productivity γ

Fixed effect regression of log labor income :

$$log(inc_{it}) = \sum_{age=20}^{65} \sum_{j \in \{G,B\}} d_t^j \cdot D_{it}^{age} \cdot \mathbf{D}_{h_{it}=j} + \gamma_i + u_{it},$$



• There are proportionately more γ_L among unheathy people.

▶ back

Results : health type prediction

| | T=5 | T=4 | T=3 | T=2 | T=1 |
|---|---|--|--|--|---|
| $h_{t_0} = P$ | 1.463 | 2.072* | 2.410 | 2.386 | 1.022 |
| $h_{t_0} = G$ | -1.457*** | -1.429*** | -1.879*** | -1.921*** | -2.250*** |
| 2^{nd} tercile of γ 3^{rd} tercile of γ | -0.247 -1.203 ^{***} | -0.337 -1.374 ^{***} | -0.509** -1.188 ^{***} | -0.546 ^{**} -1.286 ^{***} | -0.642*** -1.355 ^{***} |
| $\begin{array}{l} 2^{nd} \text{ quintile of } k_{t_0} \\ 3^{rd} \text{ quintile of } k_{t_0} \\ 4^{th} \text{ quintile of } k_{t_0} \\ 5^{th} \text{ quintile of } k_{t_0} \end{array}$ | -0.002 -0.620 -0.749 -2.348 ^{***} | -0.129 -0.429 -0.606 -1.616 ^{****} | -0.048 -0.367 -0.691* -1.169*** | -0.459 [*] -0.378 -0.701 ^{**} -1.280 ^{***} | -0.469 [*] -0.603 ^{**} -0.759 ^{***} -1.264 ^{***} |

- A lower coefficient means lower probability of worst health type (η_1)

- We control for age t₀ and cohorts (10-year bracket)



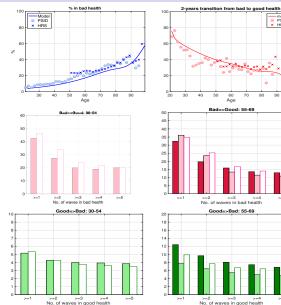
Results : Dynamics of health status: model (T=3) vs data (PSID, HRS)

PSID × HRS

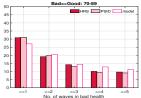
> 90 100

> > >=5

>=5





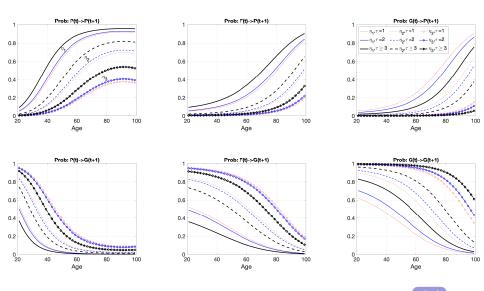




De Nardi, Pashchenko, and Porapakkarm

Lifetime Cost of Bad Health

Estimated health shock process



De Nardi, Pashchenko, and Porapakkarm

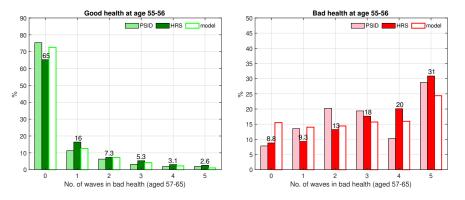
Sample from PSID: 1984-2017

| % Trans | % Transition from bad to good health conditioned on being in bad health | | | | | | | | |
|----------------------------|---|------|------|------|------|------|--|--|--|
| | >= 1 | >= 2 | >= 3 | >= 4 | >= 5 | >= 6 | | | |
| number of individual-waves | | | | | | | | | |
| 30-54 | 1420 | 646 | 375 | 230 | 141 | 83 | | | |
| 55-69 | 512 | 296 | 196 | 146 | 106 | 74 | | | |
| 70+ | 166 | 87 | 57 | 39 | 29 | 22 | | | |
| | | | | | | | | | |
| # individual | 1194 | 610 | 373 | 242 | 166 | 111 | | | |

| % Transi | % Transition from good to bad health conditioned on being in good health | | | | | | | | |
|----------------------------|--|-------|------|------|------|------|--|--|--|
| | >=1 | >= 2 | >= 3 | >= 4 | >= 5 | >= 6 | | | |
| number of individual-waves | | | | | | | | | |
| 30-54 | 11984 | 10338 | 8855 | 7461 | 6065 | 4698 | | | |
| 55-69 | 2624 | 2330 | 2113 | 1942 | 1763 | 1572 | | | |
| 70+ | 692 | 630 | 602 | 560 | 541 | 509 | | | |
| | | | | | | | | | |
| # individual | 2877 | 2554 | 2301 | 2041 | 1770 | 1505 | | | |

Distribution of unhealthy periods between 57-65: Model vs HRS

(Additional validation)



HRS: balanced panel of healthy individuals at 55



Model: working-age individuals

Consumption-saving problem

$$\max_{c_t,k_{t+1}} u(c_t,l_t,h_t) + \beta \left(\zeta_t^{\ h} E_t V_{t+1}(\mathbb{S}_{t+1}) + \left(1 - \zeta_t^{\ h}\right) \theta_{Beq} \left(\frac{k_{t+1} + k_{Beq}}{1 - \rho}\right)^{1 - \rho} \right)$$

$$\underbrace{k_t (1+r)}_{\text{total asset}} + \underbrace{exp \left(z_t^h \right) \ l_t}_{\text{labor inc}} - \text{OOP med}_{it} - \text{Ins prem} - \text{Tax} + T^{SI} (\bar{c}) = c_t + k_{t+1}$$

back

Stochastic processes estimated outside the model

Fixed effect regression of log labor income (PSID) :

$$log(inc_{it}) = \sum_{age=20}^{65} \sum_{j=\{G,B\}} d_t^j \times D_{it}^{age} \times \mathbf{D}_{h_{it}=j} + (\gamma_i + y_{it}),$$

• Health-dependent labor income process (z_t^h)

$$\begin{aligned} z_{it}^{h} &= \lambda_{t}^{h} + \gamma_{i} + y_{it} \\ y_{it} &= \rho_{y} y_{it-1} + \varepsilon_{it}; \quad \varepsilon_{it} \sim \textit{iid } N\left(0, \sigma_{\varepsilon}^{2}\right) \end{aligned}$$

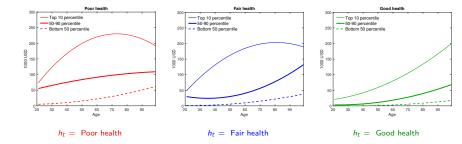
•
$$\rho_y = 0.947, \ \sigma_{\varepsilon}^2 = 0.02, \ \sigma_{z_0}^2 = 0.09, \ \sigma_{\gamma}^2 = 0.05$$

λ^h_t is used to match average labor income among worker with good, fair, and poor health

back

Health-dependent total medical expenses (x_t^h)

 \blacktriangleright x_t^h is directly estimated from MEPS

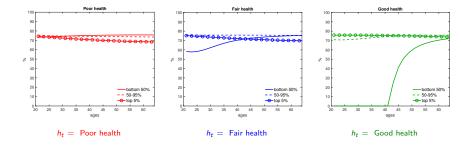


• $cvg(x_t^h, i_H)$ is estimated from people with ESI or ind insurance

• $g_t^{h,z}$ is parameterized as a logit function and estimated from MEPS

▶ back

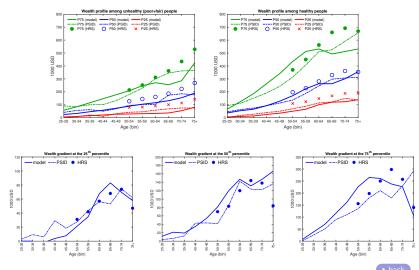
Insurance coverage : $cvg(x_t^h, i_H)$



• $cvg(x_t^h, i_H)$ is estimated from people with ESI or ind insurance

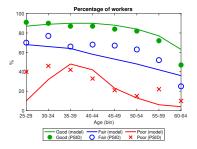
Targeted moments : model vs PSID (HRS)

Wealth health gradient

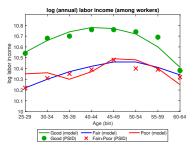


Targeted moments: Model vs PSID

Health and labor market outcomes



% Workers by health status



Average labor income (among workers) by health



Implied health gradients: Model vs PSID (HRS)

| | PSID (HRS) | | | | Model | | | |
|-------|------------|------------|-----------|------------|------------|---------|--|--|
| | bottom 1/3 | middle 1/3 | top 1/3 | bottom 1/3 | middle 1/3 | top 1/3 | | |
| 25-34 | 12% | 5% | 2% | 16% | 2% | 0% | | |
| 35-44 | 21% | 8% | 4% | 22% | 4% | 2% | | |
| 45-54 | 22% | 12% | 8% | 28% | 9% | 5% | | |
| 55-64 | 30% (36%) | 15% (20%) | 14% (13%) | 33% | 24% | 11% | | |

% unhealthy individuals in each <u>earnings tercile</u>

% unhealthy individuals in each wealth tercile

| | | PSID (HRS) | Model | | | |
|-------|------------|------------|-----------|------------|------------|---------|
| | bottom 1/3 | middle 1/3 | top 1/3 | bottom 1/3 | middle 1/3 | top 1/3 |
| 25-34 | 10% | 10% | 5% | 8% | 5% | 3% |
| 35-44 | 17% | 10% | 5% | 14% | 7% | 5% |
| 45-54 | 23% | 13% | 9% | 24% | 10% | 8% |
| 55-64 | 33% (36%) | 17% (21%) | 12% (14%) | 34% | 17% | 13% |
| 65-74 | 36% (38%) | 26% (24%) | 17% (16%) | 41% | 27% | 19% |
| 75+ | 46% (41%) | 37% (29%) | 24% (25%) | 47% | 38% | 29% |

▶ back

R1. Life-time monetary loss due to bad health (working age)

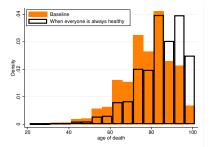
| | Over | entire life-c | ycle (21-de | ath) | Over working periods (21-64) | | | |
|--|--------------------------|-------------------------|------------------------|-----------------|------------------------------|-------------------------|------------------------|----------------|
| | All | η_1 | η_2 | η_3 | All | η_1 | η_2 | η_3 |
| % of time in bad health | 15% | 58% | 23% | 4% | 10% | 55% | 14% | 1% |
| Annual monetary losses (% of avg earning) | \$1,511 <i>(3.9%)</i> | \$8,896 <i>(23%)</i> | \$1,935 <i>(5%)</i> | \$225 (0.6%) | \$1,031 <i>(2.7%)</i> | \$7,147 <i>(18%)</i> | \$1,201 <i>(3%)</i> | \$76 (0.2%) |
| Composition (%) | | | | | | | | |
| - Medical loss paid by insurance | 36% | 33% | 39% | 39% | 32% | 33% | 33% | 18% |
| - Out-of-pocket medical loss | 27% | 22% | 30% | 36% | 20% | 20% | 21% | 11% |
| - Income losses | 37% | 45% | 31% | 24% | 48% | 47% | 46% | 71% |

- Using 2% interest rate

- Average earning (2013) is \$38,648



Variation due to age at death



| | | | When everyone are | | | |
|----------------------|------|----------|-------------------|----------|-------------------------|----------------|
| | all | η_1 | η_2 | η_3 | variation due to η | always healthy |
| Average age at death | 77.4 | 63.0 | 73.8 | 81.5 | 21 % | 83.4 |

